

Quantile Regression

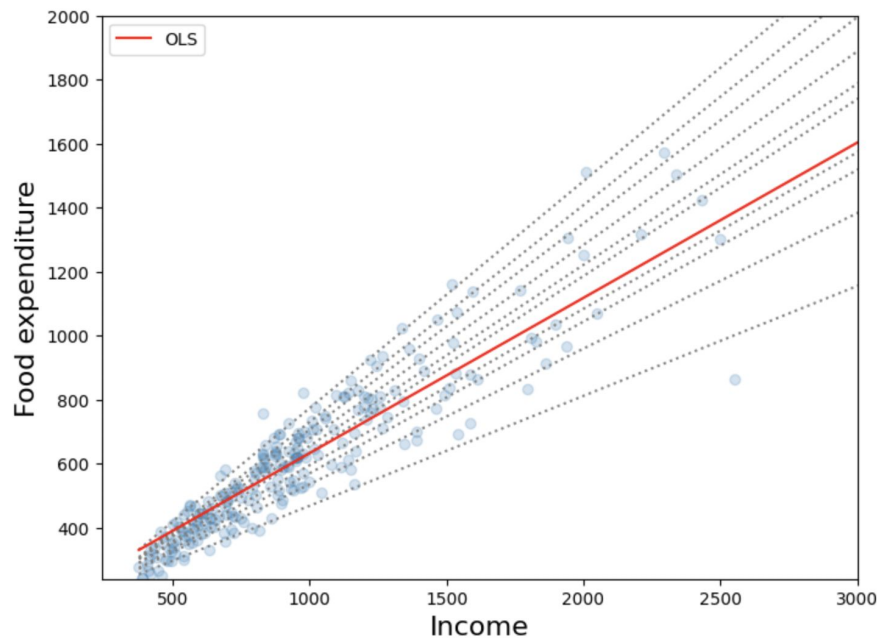
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Motivation

What the regression curve does is give a grand summary for the averages of the distributions corresponding to the set of x 's. We could go further and compute several different regression curves corresponding to the various percentage points of the distributions and thus get a more complete picture of the set.

–Mosteller and Tukey (1977)



Intuition

In OLS, we do not assume independency between predictors X and residual e . (linearity; strict exogeneity; no multicollinearity; spherical errors; normality).

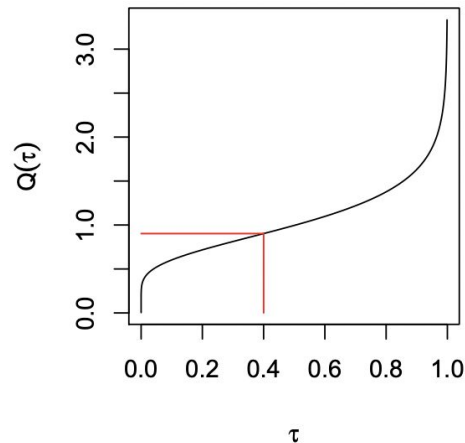
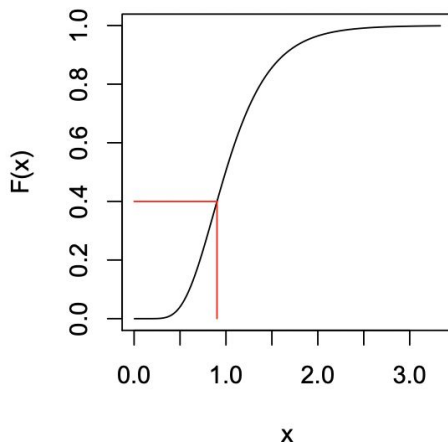
Is there a way to loosen the constraint such that the variable is independent of the error in a certain manner?

To form this question in a better way, let X be a random variable in \mathbb{R}^d , Y be a random variable in \mathbb{R} , and X, Y both have continuous distribution (density function exists). Is there a function $Q(X, t)$ such that Q is increasing in t for each X and there exists another random variable U independent from X following a uniform distribution on $[0, 1]$ that makes $Y = Q(X, U)$?

Univariate Quantile

Given a real-valued random variable X , and its cumulative distribution function F , we will define the τ th quantile of X as:

$$Q_X(\tau) = F_X^{-1}(\tau) = \inf\{x | F(x) \geq \tau\}$$



Quantile Loss Function

Optimization problem of quantile loss function:

$$\alpha_\tau = \arg \min_{\alpha} \mathbb{E}[\rho_\tau(Y - \alpha)]$$

$$\rho_\tau(u) = \tau u^+ + (1 - \tau)u^-, \quad u^+ = \max(u, 0), \quad u^- = \max(-u, 0)$$

Sketch of the proof (using FOC):

Solve $\mathbb{E}[\frac{d}{d\alpha} \rho_\tau(Y - \alpha)] = 0$.

Substitute with

$$\mathbb{E}[\frac{d}{d\alpha} (Y - \alpha)^-] = \mathbb{E}[\mathbb{1}\{Y < \alpha\}] = F_Y(\alpha)$$

Conditional Quantile

Similarly, the conditional τ th quantile solves:

$$\hat{\alpha}_\tau(x) = \arg \min_{\alpha} \mathbb{E}[\rho_\tau(Y - \alpha) | X = x]$$

Assume linear formation, we have the following parametric form

$$\hat{\beta}_\tau(x) = \arg \min_{\beta} \mathbb{E}[\rho_\tau(Y - X^T \beta) | X = x]$$

$$\begin{aligned} Q_{Y|X}(\tau|x) &= \inf\{y : F(y) \geq \tau | X = x\} \\ &= \sum_k \beta_k(\tau) x_k = \beta(\tau)^T x \end{aligned}$$

Back to Intuition – Specified Quantile Regression

Is there a function $Q(X, t)$ such that Q is increasing in t for each X and there exists another random variable U independent from X following a uniform distribution on $[0, 1]$ that makes $Y = Q(X, U)$? The goal is to correctly define U .

Define U such that

$$Y = Q_{Y|X}(U|X)$$

Recall the inverse relationship

$$\begin{aligned}\mathbb{P}(U < t \mid X = x) &= \mathbb{P}(F_{Y|X}(Y|x) < t \mid X = x) \\ &= \mathbb{P}(Y < Q_{Y|X}(t|x) \mid X = x) \\ &= t\end{aligned}$$

Hence, we obtain a uniform distribution fully independent from X .

$$U = F_{Y|X}(Y|X) = \mu[0, 1]$$

Regression Quantiles Computation: Linear Programming

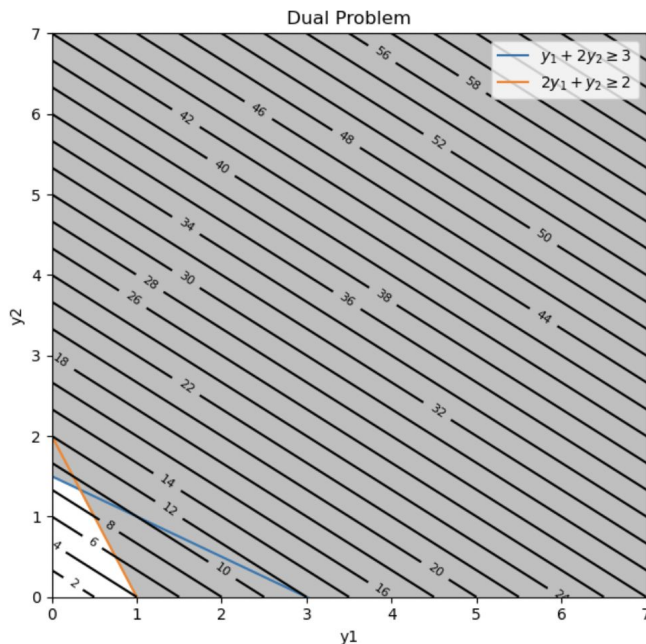
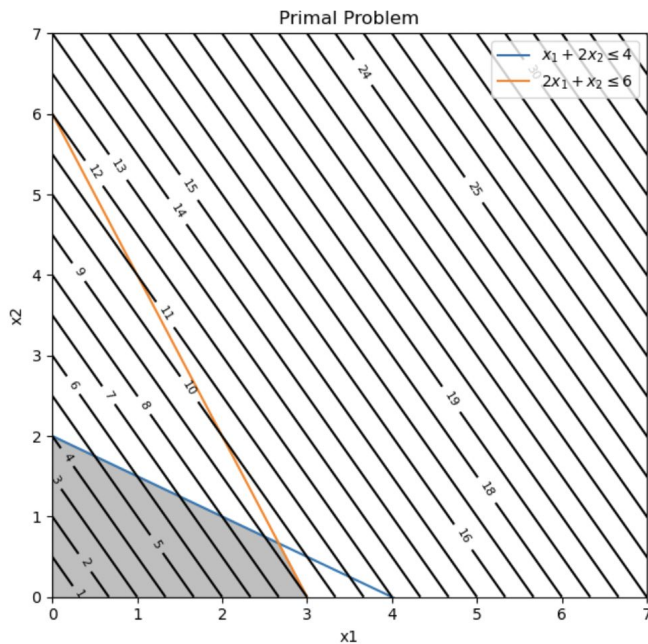
For every linear programming problem (the primal problem), there exists an associated LP problem called its dual problem (there's complementary slackness.)

Primal:

$$\begin{aligned} &\text{maximize} && c^T x \\ &\text{subject to} && Ax \leq b \\ &&& x \geq 0 \end{aligned}$$

Dual:

$$\begin{aligned} &\text{minimize} && b^T y \\ &\text{subject to} && A^T y \geq c \\ &&& y \geq 0 \end{aligned}$$



Linear Programming - Continued

The goal is to minimize

$$\mathbb{E}[\rho_\tau(Y - X^\top \beta)]$$

We consider its sample version for the ease of coding

$$\arg \min_{\beta \in \mathbb{R}^k} \sum_{i=1}^n \rho_\tau(y_i - x_i^\top \beta)$$

Let $u_i = (y_i - x_i^\top \beta)^+$ and $v_i = (y_i - x_i^\top \beta)^-$ with $u_i, v_i \geq 0$. $u_i - v_i = y_i - x_i^\top \beta$.

$$\begin{aligned} \arg \min_{\beta \in \mathbb{R}^k} \sum_{i=1}^n \rho_\tau(y_i - x_i^\top \beta) &= \arg \min_{\beta \in \mathbb{R}^k} \sum_{i=1}^n \rho_\tau(u_i - v_i) \\ &= \arg \min_{\beta \in \mathbb{R}^k} \sum_{i=1}^n \tau u_i + (1 - \tau) v_i \\ &= \arg \min_{\beta \in \mathbb{R}^k} [u^\top \mathbf{1} \tau + v^\top \mathbf{1} (1 - \tau)] \end{aligned}$$

Primal - Dual Problem

Original:

$$\arg \min_{\beta \in \mathbb{R}^k} [u^\top \mathbf{1} \tau + v^\top \mathbf{1} (1 - \tau)]$$

$$s.t. \ y - X^\top \beta - (u - v) = 0, \ (u, v \geq 0)$$

After Lots of Linear Algebra...

Primal: Let $w = (\beta, u, v)$, $c = (0, \mathbf{1} \tau, \mathbf{1} (1 - \tau))$, $Z = (X^\top, I_n, -I_n)$

$$\arg \min_{w \in \mathbb{R}^k} (w^\top c)$$

$$s.t. \ Z w = y$$

Dual:

$$\arg \max_z (y^\top z)$$

$$s.t. \ X z = (1 - \tau) X \mathbf{1}, z \in [0, 1]^n$$

Another Way

$$\mathbb{E}[\tau P + (1 - \tau)N]$$

$$s.t. \quad P - N = Y - X^\top \beta$$

where $P = (Y - X^\top \beta)^+$ and $N = (Y - X^\top \beta)^-$ Call - Put Parity

We may rewrite and eliminate N to get the dual problem

$$\min_{P \geq 0, \beta} \mathbb{E}[P + (1 - \tau)X^\top \beta]$$

$$s.t. \quad P + X^\top \beta \geq Y$$

Add a slack variable V to the dual problem

$$\min_{P \geq 0, \beta} \mathbb{E}[P + (1 - \tau)X^\top \beta] + \max_{V \geq 0} [V(Y - P - X^\top \beta)]$$

We may rewrite it by combining the min max

$$V_D = \min_{P \geq 0, \beta} \max_{V \geq 0} \mathbb{E}[P + (1 - \tau)X^\top \beta + V(Y - P - X^\top \beta)]$$

Using the minimax inequality we have $V_P \geq V_D$ with

$$\begin{aligned} V_P &= \max_{V \geq 0} \min_{P \geq 0, \beta} \mathbb{E}[P + (1 - \tau)X^\top \beta + V(Y - P - X^\top \beta)] \\ &= \max_{V \geq 0} \mathbb{E}[VY] + \min_{P \geq 0, \beta} \mathbb{E}[(1 - V)P + (1 - \tau - V)X^\top \beta + VY] \end{aligned}$$

Primal \rightarrow

Then, we have obtained the primal problem

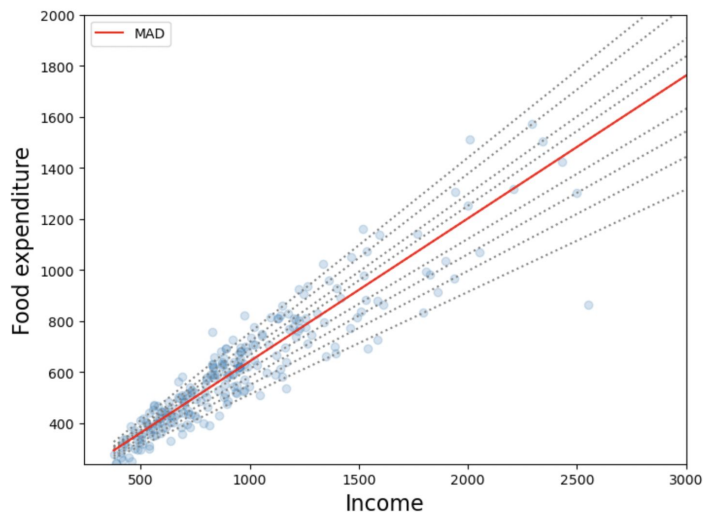
$$\max_{V \geq 0} \mathbb{E}[YV]$$

$$s.t. \quad V \leq 1 \quad [P_t \geq 0]$$

$$\mathbb{E}[VX] = (1 - \tau)\mathbb{E}[X]$$

Realization

Solve the LP problem
using Gurobi



```
income = np.array(engle_data['income'])
food = np.array(engle_data['food'])
housing = np.array(engle_data['housing'])
nbi=len(income)
X_i_k = np.array([np.ones(nbi),income]).T
Y = np.array([food,housing]).T
_,nbk = X_i_k.shape
qr_lp=grb.Model()
τ = 0.5
P = qr_lp.addMVar(shape=nbi, name="P")
β = qr_lp.addMVar(shape=nbk, name="β", lb=-grb.GRB.INFINITY)
qr_lp.setObjective(np.ones(nbi) @ P + (1-τ) * (np.ones(nbi) @ X_i_k) @ β, grb.GRB.MINIMIZE)
qr_lp.addConstr(P + X_i_k @ β >= food)
qr_lp.optimize()

βhat = qr_lp.getAttr('x')[-nbk:]
βhat
```

Gurobi Optimizer version 10.0.2 build v10.0.2rc0 (mac64[arm])

CPU model: Apple M1

Thread count: 8 physical cores, 8 logical processors, using up to 8 threads

Optimize a model with 235 rows, 237 columns and 705 nonzeros

Model fingerprint: 0x2e1ba335

Coefficient statistics:

Matrix range [1e+00, 5e+03]

Objective range [1e+00, 1e+05]

Bounds range [0e+00, 0e+00]

RHS range [2e+02, 2e+03]

Presolve time: 0.00s

Presolved: 235 rows, 237 columns, 705 nonzeros

Iteration	Objective	Primal Inf.	Dual Inf.	Time
0	handle free variables			0s
223	8.2118037e+04	0.000000e+00	0.000000e+00	0s

Solved in 223 iterations and 0.01 seconds (0.00 work units)

Optimal objective 8.211803659e+04

[81.48614818083601, 0.560174688436782]

Compare the beta values

```
In [17]: import statsmodels.api as sm
import statsmodels.formula.api as smf
import matplotlib.pyplot as plt
import scipy.sparse as spr
```

```
In [18]:  $\tau = 0.5$ 
model = smf.quantreg('food ~ income', engle_data)
print(model.fit(q= $\tau$ ).summary())
```

```
QuantReg Regression Results
=====
Dep. Variable:          food    Pseudo R-squared:          0.6205
Model:                  QuantReg    Bandwidth:              64.51
Method:                 Least Squares    Sparsity:              209.3
Date:                   Thu, 13 Jul 2023    No. Observations:      235
Time:                   15:40:38    Df Residuals:          233
                                   Df Model:              1
=====

```

	coef	std err	t	P> t	[0.025	0.975]
Intercept	81.4862	14.635	5.568	0.000	52.652	110.321
income	0.5602	0.013	42.513	0.000	0.534	0.586

```
=====
```

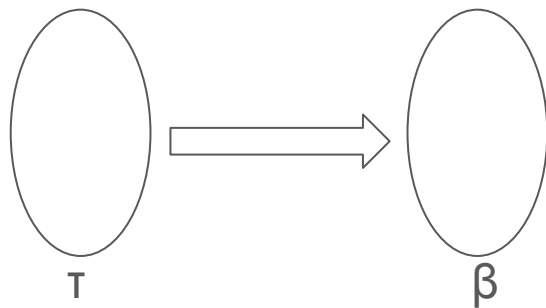
The condition number is large, 2.38e+03. This might indicate that there are strong multicollinearity or other numerical problems.

Quantile Curve

The previous optimization problem has provided a way to compute β for pointwise values of τ .

If we want to compute the whole curve that maps τ to β , the pointwise method surely makes it an impossible task to finish...

The solution is not difficult, we “integrate” the infinite amount of optimization problem into one!



Continued

Since we know that β_τ solves the primal problem

$$\max_{V_{\tau_i} \in [0,1]} \mathbb{E}[Y V_{\tau_i}]$$

$$s.t. \mathbb{E}[V_{\tau_i} X] = (1 - \tau_i) \mathbb{E}[X] \quad [\beta]$$

for $\tau_i = \frac{i}{n}$, $0 \leq i \leq n$

Combine them by taking the sum of these problems,

$$\max_{V_{\tau_i} \in [0,1]} \sum_{\tau_i} \mathbb{E}[Y V_{\tau_i}]$$

$$s.t. \forall i, \mathbb{E}[V_{\tau_i} X] = (1 - \tau_i) \mathbb{E}[X]$$

Taking the limit and we may convert the sum into an integral

$$\max_{V(\cdot) \geq 0} \int_0^1 \mathbb{E}[V_\tau Y] d\tau$$

$$s.t. V(\tau) \leq 1$$

$$\mathbb{E}[V(\tau) X] = (1 - \tau) \mathbb{E}[X]$$

Similarly, its dual is constructed as

$$\min_{P \geq 0, \beta} \int_0^1 \mathbb{E}[P(\tau) + (1 - \tau) X^T \beta(\tau)] d\tau$$

$$s.t. P(\tau) \geq Y - X^T \beta(\tau)$$

Primal \rightarrow

Dual \rightarrow

Monotonicity Constraint

Recall the definition of quantile as the inverse of a CDF function. There is a natural constraint of monotonicity imposed on the dual (Koenker and Ng):

$$\begin{aligned} \min_{P \geq 0, N \geq 0, \beta} & \int_0^1 \mathbb{E} [P(\tau) + (1 - \tau)X^\top \beta(\tau)] d\tau \\ \text{s.t. } & P(\tau) - N(\tau) = Y - X^\top \beta(\tau) \\ & X^\top \beta(\tau) \geq X^\top \beta(\tau'), \tau \geq \tau' \end{aligned}$$

To solve it more easily, we consider its primal formulation (Carlier, Chernozhukov and Galichon). Given that

$$V(\tau) = 1 \{Y \geq X^\top \beta(\tau)\},$$

We have $X^\top \beta(\tau)$ nondecreasing in $\tau \implies V(\tau)$ nonincreasing.
The primal problem is

$$\begin{aligned} \max_{V(\tau)} & \int_0^1 \mathbb{E}[YV(\tau)]d\tau \\ \text{s.t. } & 0 \leq V(\tau) \leq 1 \\ & \mathbb{E}[V(\tau)X] = (1 - \tau)\mathbb{E}[X] \\ & V(\tau) \leq V(\tau'), \tau \geq \tau' \end{aligned}$$

Sampled Version

Assume $\tau_1 = 0 < \dots < \tau_T \leq 1$ and let \bar{x} be a $1 \times K$ row vector whose k -th entry is $\mathbb{E}[X_k]$. I is the dimension of Y . T is the size of the partition set of τ . The sampled version of the previous primal problem is

$$\begin{aligned} & \max_{V_{ti} \geq 0} \frac{1}{I} \sum_{\substack{1 \leq i \leq I \\ 1 \leq t \leq T}} V_{ti} Y_i \\ & s.t. \quad V_{ti} \leq 1 \\ & \quad \frac{1}{I} \sum_{1 \leq i \leq I} V_{ti} X_{ik} = (1 - \tau_t) \bar{x}_k \\ & \quad V_{(t+1)i} \leq V_{ti} \end{aligned}$$

Since $\tau_1 = 0$, we have $V_{1i} = 1$ from the second constraint. The program becomes

$$\begin{aligned} & \max_{V_{ti}} \frac{1}{I} \sum_{\substack{1 \leq i \leq I \\ 1 \leq t \leq T}} V_{ti} Y_i \\ & s.t. \quad V_{1i} = 1 \\ & \quad \frac{1}{I} \sum_{1 \leq i \leq I} V_{ti} X_{ik} = (1 - \tau_t) \bar{x}_k \\ & \quad V_{1i} \geq V_{2i} \geq \dots \geq V_{(T-1)i} \geq V_{Ti} \geq 0 \end{aligned}$$

Next, let τ be the $T \times 1$ row matrix with entries τ_k and D be a $T \times T$ matrix defined as

$$D = \begin{pmatrix} 1 & 0 & 0 & \cdots & 0 & 0 \\ -1 & 1 & 0 & \ddots & \vdots & \vdots \\ 0 & -1 & 1 & \ddots & 0 & 0 \\ \vdots & \ddots & \ddots & \ddots & 0 & 0 \\ \vdots & & 0 & -1 & 1 & 0 \\ 0 & & 0 & 0 & -1 & 1 \end{pmatrix}$$

we have $V^\top D \geq 0$ if and only if

$$V_{1i} \geq V_{2i} \geq \dots \geq V_{(T-1)i} \geq V_{Ti} \geq 0$$

← replace with D

We can write the sampled version into the following matrix form

$$\begin{aligned} & \frac{1}{I} \max_V 1_T^\top VY \\ \text{s.t. } & \frac{1}{I} VX = (1_T - \tau) \bar{x} \\ & V^\top D 1_T = 1_I \\ & V^\top D \geq 0 \end{aligned}$$

Suppose $\pi = \frac{D^\top V}{I}$ and $U = D^{-1} \mathbf{1}_I$, $\mu = D^\top (1_T - \tau)$ and $p = \frac{1_T}{I}$. It is equivalent to

$$\begin{aligned} & \max_{\pi} U^\top \pi Y \\ \text{s.t. } & \pi X = \mu \bar{x} \\ & \pi^\top 1_T = p \\ & \pi \geq 0 \end{aligned}$$

Assume that the first entry of X is one for the ease of computation. If π satisfies the constraints

$$\sum_{i=1}^I \pi_{ti} = \mu_t \text{ and } \sum_{t=1}^T \pi_{ti} = p_i$$

then π can be thought of as a joint probability on τ and X given the marginal probability constructed above.

After all these
matrix algebra...

We defined a joint
probability density,
 π !

1D Vector Quantile Regression

1D-VQR is equivalent to classical quantile regression by construction!

It can be rewritten into the one-dimensional vector quantile regression construction in the continuous case (Carlier, Chernozhukov and Galichon)

$$\begin{aligned} & \max_{\pi} \mathbb{E}_{\pi}[UY] \\ & s.t. \ U \sim \mu \\ & \quad (X, Y) \sim P \\ & \quad \mathbb{E}[X | U] = \mathbb{E}[X] \end{aligned}$$

If we replace the mean-independence between X and U by independence using conditional probability (scalar VQR, by *Thm.3.3*), we have

$$\begin{aligned} & \max_{\pi} \mathbb{E}_{\pi}[UY] \\ & s.t. \ U \sim \mu \\ & \quad (X, Y) \sim P \\ & \quad X \perp\!\!\!\perp U \end{aligned}$$

The solution to the latter problem is simply $U = F_{Y|X}(Y | X)$ and the nonparametric conditional quantile representation is $Y = F_{Y|X}^{-1}(U | X)$. This conforms with the classical quantile function we constructed above!

Computation

```
In [21]: D_t_t = spr.diags([1, -1], [ 0, -1], shape=(nbt, nbt))

U_t_1 = np.linalg.inv(D_t_t.toarray()) @ np.ones( (nbt,1))
μ_t_1 = D_t_t.T @ (np.ones((nbt,1)) - τ_t_1)

A1 = spr.kron(spr.identity(nbt),X_i_k.T)
A2 = spr.kron(np.array(np.repeat(1,nbt)),spr.identity(nbi))
A = spr.vstack([A1, A2])
rhs = np.concatenate( [(μ_t_1 * xbar_1_k).flatten(), np.ones(nbi)/nbi])
obj = np.kron(U_t_1, Y_i_1).T
vqr_lp=grb.Model()
pi = vqr_lp.addMVar(shape=nbi*nbt, name="pi")
vqr_lp.setParam( 'OutputFlag', False )
vqr_lp.setObjective( obj @ pi, grb.GRB.MAXIMIZE)
vqr_lp.addConstr(A @ pi == rhs)
vqr_lp.optimize()

φ_t_k = np.array(vqr_lp.getAttr('pi'))[0:(nbt*nbk)].reshape((nbt,nbk))

βvqr_t_k = D_t_t.toarray() @ φ_t_k

βvqr_t_k[10,:]
```

```
Out[21]: array([81.48614818,  0.56017469])
```

$$\max_{\pi} U^{\top} \pi Y$$

$$s.t. \pi X = \mu \bar{x}$$

$$\pi^{\top} 1_T = p$$

$$\pi \geq 0$$

Vectorize

$$\begin{pmatrix} I_T \otimes X^{\top} \\ 1_T^{\top} \otimes I_I \end{pmatrix} \text{vec}(\pi) = \begin{pmatrix} \text{vec}(\mu \bar{x}) \\ p \end{pmatrix}$$

Vector Quantile Regression - General Case

In some fixed nonatomic probability space, $(\Omega, \mathcal{F}, \mathbb{P})$, given a random vector Z with values in \mathbb{R}^k defined on this space, we will denote by $\mathcal{L}(Y)$ the law of Z . We fix as a reference measure the uniform measure on the unit cube $[0, 1]^d$

$$\mu_d := \mathcal{U}([0, 1]^d)$$

Given Y , an integrable \mathbb{R}^d -valued random variable on $(\Omega, \mathcal{F}, \mathbb{P})$, Brenier's Theorem states that there exists a unique $U \sim \mu_d$ and a unique convex function defined on $[0, 1]^d$ such that

$$Y = \nabla \varphi(U).$$

The map $\nabla \varphi$ is called the **Brenier's map** between μ_d and $\mathcal{L}(Y)$.

The **vector quantile** of Y is defined to be the Brenier's map between μ_d and $\mathcal{L}(Y)$.

In one-dimensional space, the optimal transport map of Brenier is given by $\nabla \varphi = Q$, where Q is the quantile of Y . Monotonicity persists in both one-dimensional and higher dimensions.

Brenier's Continued

Theorem 2.1. (*Brenier's theorem*) *If Y is a squared-integrable random vector valued in \mathbb{R}^d , there is a unique map of the form $T = \nabla\varphi$ with φ convex on $[0, 1]^d$ such that $\nabla\varphi_{\#}\mu = \text{Law}(Y)$, this map is by definition the vector quantile function of Y .*

More topics to explore:

Regularized Vector Quantile Regression

From Quasi-specified QR to Unspecified QR

...

Reference

Carlier, G., Chernozhukov, V., De Bie, G., Galichon, A. (2020). *Vector Quantile Regression and Optimal Transport, from Theory to Numerics*. Empirical Economics. <https://doi.org/10.48550/arXiv.2102.12809>

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